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### THE SKEW FREQUENCY DISTRIBUTIONS AND THE FUNDAMENTAL LAW OF THE GEOCHEMICAL PROCESSES<sup>1</sup>

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#### ABSTRACT

The frequency distribution of concentrations of chemical elements has been investigated. Modern statistical methods for the analysis of large and small samples have been used. As a result of the work, the formulation of the fundamental law of the geochemical processes is proposed. This law can be applied in many cases to the investigation of deposits. The author points out the difference between "the distribution of the concentrations" in V. M. Goldschmidt's sense and "the probability distribution of the concentrations" of the present paper.

#### INTRODUCTION

One of the most valid problems of geochemistry is the investigation of the frequency distribution of the concentration of elements in igneous, metamorphic, and sedimentary rocks, minerals, soils, and organic matter. The positive solution of these problems will give the following: (a) a check of theoretical schemes of geochemistry by comparison of the theoretical probability distribution with empirical frequency distribution and (b) mean values of the concentrations of the elements and quantitative measurement of the correlation between the members of paragenesis estimations.

Richardson's paper (Richardson and Sneesby, 1922) is the first stage in the study of the frequency distribution in geochemistry, as it is known by the author. This paper won the broad approval of the petrologists and was theoretically interpreted by Richardson himself (1923). So in petrochemistry the theoretical sense of the frequency distributions was pointed out at the very beginning. Only five years before

Richardson's paper was published, there had appeared A. A. Markov's mathematical proof of the physical sense of the probability distribution (Markov, 1917). Markov's very valuable paper aroused much interest in the mathematical literature, but these works remained completely unknown to geologists.

Since Richardson's paper was published, we have not had any frequency distribution investigation in geochemistry for seventeen years. But during these years F. J. Loevinson-Lessing published many papers on the frequency distributions of his "magmatic coefficients." These papers were stimulating to many Russian petrologists and to the author in particular.

In 1940 N. K. Rasumovsky wrote about the wide expansion of lognormal frequency distribution and particularly about the frequency distribution of the concentrations of elements. The wide expansion of lognormal distribution pointed out by Rasumovsky was taken into account by A. N. Kolmogorov, who gave his stochastic schemes by using modern methods (Kolmogorov, 1941). In 1948 Rasumovsky returned to his idea of lognormal distribution (Rasumovsky, 1948).

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During this period the lognormal distribution has been applied in the prospecting of gold.

Systematic analysis of frequency distributions for geochemical purposes investigated by the author have been carried out since 1943. In the course of time he studied the geochemical sense of the frequency distribution of the coefficients of porosity in Upper Permian beds of Buguruslan (Vistelius, 1945), the frequency distribution of six-valence sulfur in Paleozoic beds between the River Volga and the Ural Moun-(Vistelius, 1949). The stochastic foundation (Vistelius and Sarmanov, 1947) and statistical experimentation were presented in these papers as well as the application of the tests of the goodness of fit between the calculated functions and the observed frequencies. The basic conclusions of the author are as follows: the frequency distribution reflects the geochemical environment of sedimentation. The validity of the investigation of the frequency distributions was also pointed out by the author in some papers published later (Vistelius, 1948, 1952).

L. H. Ahrens' first paper (1953) on frequency distributions proposed the lognormal type of distribution and formulated his fundamental law of geochemistry. Three papers on this question were published by Ahrens later (1954a, 1954b, 1957). The whole scope of Ahrens' papers aroused the great interest of many geochemists in frequency distributions. His ideas gave rise to some critical articles (Aubrey, 1954, 1956; Chayes, 1954; Miller and Goldberg, 1955; Durovič, 1957). In Ahrens' papers the facts connected with frequency distributions were given for the first time as the material for the formulation of the law of geochemistry.

The importance of Ahrens' propositions, the disputability of them, and the many aspects of the problems and methods create a confused situation.

The author of this paper wished to avoid polemics and therefore preferred to analyze the frequency distributions but not to analyze general considerations given previously.

# PRINCIPLES OF FREQUENCY ANALYSIS AND CHECK OF PREVIOUSLY PUBLISHED DATA

The frequency distribution analysis is a means of checking the hypothesis on the mechanism of the geochemical process which takes place for the chemical element x with the function of the probability distribution, p(x). We formulate the problem as follows: The theory of the geochemical process concerning the element x gives the hypothesis  $H_0$ . This hypothesis is called the "null hypothesis  $H_0$ ." We deduced from  $H_0$  the probability distribution function of p(x). Observations of the values of x as  $x_0, x_1, \ldots, x_n$  give the frequencies f(x). If the same function of difference,

$$\delta = F[f_i(x) - p_i(x)], \qquad (1)$$

is small for the points i, we say that  $H_0$  is not contradictory to our observations. If  $\delta$  is large, we say that  $H_0$  does not satisfy our observations and must be rejected as false.

In routine investigations we come across the following cases.

- a)  $H_0$  and consequently p(x) are theoretically deduced from concrete investigations of the analyzed object, i.e., from the scheme of a geochemical process. In this case, in comparing p(x) and f(x), we can analyze the concrete sense of  $\delta$  as well as that of f(x). In other words, even increased values of  $\delta$  can in some cases be explained and  $H_0$  accepted.
- b)  $H_0$  and consequently p(x) are not theoretically deduced from the geochemical sense of the process but are the explanation of frequencies f(x) which are purely empirically smoothed by p(x). In this case there are very severe restrictions on the value of  $\delta$ . If these conditions are satisfied, we say that  $H_0$  coincides with our observations (they coincide but are not proved!). If our conditions are not satisfied,  $H_0$  is false.

In some cases we cannot say whether  $H_0$  coincides or not with our observations; the values of  $\delta$  are at the boundary of the zones

of the rejection and acceptance of  $H_0$ . These conditions are in some way or other based on a number of observations.

#### INVESTIGATION OF THE PUBLISHED DATA

We shall study in detail the data published by N. K. Rasumovsky (1940) and L. H. Ahrens (1954a, b, 1957). The data of other authors (Aubrey, 1954; Durovič, 1957) have been published in the form of diagrams. The very small scale of these diagrams prevents us from analyzing them.

All the analyzed data may be divided into four groups.

The first group is given as the frequency distributions of Ga(g), Sc(d), K(g), Rb(g), Zn(g), V(d), La(g), Sc(g), Pb(g), V(g), and Cr(g).<sup>2</sup>

These distributions are given in the form of histograms by Ahrens (1954a, b, 1957). The cited histograms have only three classes with more than five observations in each class. Other classes have less than five observations in every class. According to statistics, we know that the goodness of fit of the calculated functions and the observed frequencies must be performed when every class has had more than five observations. Consequently, after excluding the classes with less than five observations, we obtain histograms with three classes only. But the normal distribution has three parameters; therefore, the number of degrees of freedom equals the quantity of classes minus three; as we have only three classes, we must have the highest goodness of fit with a calculated normal function with any three unequal numbers. So the first group of the data published cannot be examined.

The second group of the published observations contains histograms with frequencies having more observations than five in each class and with the number of classes more than three, Here are the data of Ahrens (1954a, b, 1957) on Cs(g), Mo(g), Ga(d), Co(d), Zr(d), Be(g), K(d), Rb(d), F(g), Mo in igneous rocks, Li in muscovite,

 $^2$  Cr(g) is chromium in granite; Sc(d) is scandium in diabase, etc.

and Rb in biotite, and of Rasumovsky (1940) on Cu, Pb, and Zn. These histograms have one or more degrees of freedom and can be investigated in some detail.

The smoothing of the observed frequencies in equal logarithmic values of classes by means of a normal curve for this distribution has been produced. After having calculated the ordinates of a normal curve, we compare them with the observed frequencies. For the sake of checking the goodness of fit of the calculated and observed data, we shall take Pearson's

$$\chi^{2} = \sum_{i} \frac{[f(x_{i}) - p(x_{i})]^{2}}{p(x_{i})}, \quad (2)$$

where  $\chi^2$  is the test of the goodness of fit,  $f(x_i)$  is the observed and  $p(x_i)$  the calculated frequencies by the lognormal  $H_0$ . As we know, the distribution of  $\chi^2$  with given degrees of freedom gives the level of confidence of  $H_0$  in percentage. For the evaluation of the level of confidence, many tables were published which have at the head of the rows the degrees of freedom, whereas at the head of the columns are the levels of confidence, and within the table are the figures of  $\chi^2$ . If we compare the calculated and observed frequencies without any theoretical background of  $H_0$  (as it is in our case), we must take such great values of confidence as 50 per cent or more. Values of confidence smaller than 50 per cent give indefinite results of investigations; we can neither reject  $H_0$  nor accept it.3

In table 1 observed and calculated frequencies, degrees of freedom,  $\chi^2$ , and levels of confidence are given.

The technique of calculations of the nor-

 $^3$  The value of the confidence level cannot be obtained in a pure theoretical way. As a rule, it must be revealed by experience. We have no data in our case about the mutual independence of observations. In such cases it is necessary to be very cautious in the acceptance of a new hypothesis. The above points out why we have accepted  $H_0$  with the confidence level higher than 50 per cent and we suppose the 15 per cent level insufficient for acceptance. In our case the common confidence level used for independent observations and equal to 5 per cent cannot be used for acceptance of  $H_0$ .

mal curve by means of the observed  $\bar{x}$  and S are illustrated in table 2. By repeating calculations similar to those of table 2, the reader can check all the other data in table 1.

The data from table 1 tell us that the lognormal  $H_0$  can be more or less accepted for Cu in ores of Altai, Rb in biotite, Li in muscovite, K in diabase, Mo and Be in

granites. The lognormal  $H_0$  must be rejected for molybdenium in igneous rocks, Zr in diabase, and Pb in the ores of Altai. We must also reject the lognormal  $H_0$  for As in granites and Cr in diabase. The geometrical peculiarities of these histograms (Ahrens, 1954, 1957) have no features of a normal curve, which is obvious without

TABLE 1 FREQUENCY DISTRIBUTIONS OF SOME CHEMICAL ELEMENTS (y Is Observed and y' Is Calculated Frequencies) (Lognormal  $H_0$ )

Elements Studied	Observed (y) and Calcu- LATED (y')			F	REQUEN	CIES OF	тне С	ASSES				No. of Obser-	χ²	Degrees or Free-	Confidence Coefficients
*******	FRE- QUENCIES	1	2	3	4	5	6	7	8	9	10	VATIONS		DOM (k)	
(L. H. AHRENS)  K in diabases	\[ \begin{cases} y \\ y' \end{cases} \]	3_9	7	9	12 10.7	5 9	4					40	0.51	1	$0.50 > P_k > 0.30$
Rb in diabases	\\ \{ \bar{y}{y'} \end{array}	$\phantom{aaaaaaaaaaaaaaaaaaaaaaaaaaaaaaaaaaa$	7	3	7	8 6.7	4 6	3				34	2.81	1	$0.10 > P_k > 0.05$
F in granites	y y'	1_8	7	15 14.7	20 16.2	7	3	1				54	1.73	1	$0.20 > P_k > 0.10$
Mo in igneous rocks	\\ \begin{cases} y \\ y' \end{cases}	1_4	5	6 14.3	35 26.2	24 23.5		1	1			81	9.03	2	$0.02 > P_k > 0.01$
Ga in diabases	y y'	5 7.4	26 19.7	17 20.9	7 9	2						57	3.52	1	$0.10 > P_k > 0.05$
Co in diabases	y y'	2	1 1 11.2	5	18 17.5	22 17.4	8	9				57	2.47	1	$0.20 > P_k > 0.10$
Zr in diabases	y y'	2 6	6	7 13.9	25 17.4	9	3					52	7.43	1	$0.01 > P_k > 0.00$
Rb in biotites		15 12.6	30 32.8	51 53.7	51 46.2	20 21.0	5 5.7					172	0.98	3	$0.90 > P_k > 0.80$
Cs in granites	\\ \{ \bar{y}' \\ \p' align*	15	4 5.1	9 6.1	$\underbrace{\frac{3}{12}}$	.8	4_4	1 0				28	2.76	1	$0.10 > P_k > 0.08$
Mo in granites	\ \{ \bar{y}{y'} \}	1_5	5	8 11.4	22 17.5	14 15.7	9_	11.0				61	2.50	2	$0.30 > P_k > 0.20$
Be in granites	$\left\{\begin{matrix} y \\ y' \end{matrix}\right.$	1_4	1.5	6 8.1	16 12.1	9 11.6	6_10	5				47	2.44	2	$0.30 > P_k > 0.20$
Li in muscovites.	$\begin{cases} y \\ y' \end{cases}$	6 7.0	12 10.9	22 17.0	16 19.0	12 15.0	5 13	9				82	2.8	3	$0.50 > P_k > 0.30$
(N.K.RASUMOVSKY)	(y	13 9.0	33 42.0	125 128.1	242 231.0	247 245.9	138 154.4	83	1 6			882	8.20	4	$0.10 > P_k > 0.00$
Рь	y y'	1 11	9 2.9	36 32.4	100 76.2	111 133.7	148 173.6	178 168.0	123 120.3	91 64.2	19 34.7	816	35.2	6	$0.001 > P_k$
Cu	$\begin{cases} y \\ y' \end{cases}$	2 23	25	65 66.6	127 142.5	198 189.2	168 156.0	76 79.7	24	30.8	1	688	4.4	0 4	$0.50 > P_k > 0.3$

TABLE 2 Example of Calculations of Normal Curve by Means of the Observed  $\bar{x}$  and S (Equal Logarithmic Classes)

Logarithmic Classes	No. of Classes (x)	Frequencies Observed (y)	$x-\bar{x}$	$t = (x - \bar{x})/s$	$\phi(t_i)$	$\phi(t_i) - \phi(t_{i-1})$	y <b>'</b>	$\Delta =  y-y' $	$\Delta^2$	$\Delta z/y'$
0.01- 0.03 0.03- 0.06	1 2	$\binom{2}{25}$ 27	- ∾ -3.594 -2.594	$-\infty$ $-2.535$ $-1.829$	-0.5000 -0.4944 -0.4663	0.0056 0.0281	3.85 19.33 23.18	3.82	14.59	0.75
0.60- 0.125 0.125-0.250	3 4	65 127	-1.594	-1.124	-0.3695	0.0968 0.2071	66.60	1.60 15.48	2.56	0.04 1.68
0.250-0.500	5	198	-0.594 +0.406	-0.419 +0.286	-0.1624 +0.1126	0.2750	189.21	8.79	77.26	0.41
1.000-2.000	6 7	168 76	+1.406 +2.406	+0.992 +1.697	+0.3394 +0.4552	0.2268 0.1158	156.0 <del>4</del> 79.67	11.96 3.67	143.04 13.47	0.92 0.17
2.000-4.000 4.000-8.000	8 9	$\begin{bmatrix} 24 \\ 2 \\ 27 \end{bmatrix}$	+3.406	+2.402	+0.4918	0.0366 0.0073	25.18 5.02 30.82	3.82	14.59	0.47
3.000-16.000	10	1)	+4.406 +∾	+3.107 +∞	+0.4991 +0.5000	0.0009	0.62			
Total		688				1.0000	688.00			4.44

 $\bar{x} \approx 5.094$ , S = 1.418,  $\chi^2 = 4.44$ , k = 4,  $0.5 > P_k(\chi^2) > 0.3$ 

making any calculations. All the other distributions given by Ahrens (1954a, b) and the distribution of Zn given by Rasumovsky have such frequencies that they do not permit acceptance or rejection of the lognormal  $H_0$ . These data are quite indefinite (for example, the frequency distribution of Zn has a confidence coefficient between 5 and 10 per cent).

The third group of observations contains Ahrens' numerical data which are given in table 1 of this paper (Ahrens, 1954a). These data cannot be analyzed by a frequency grouping as Ahrens (1954a) did because the number of his observations is too small. But they can be analyzed by one of the newest methods, which we shall give later (Dunin-Barkowsky and Smirnoy, 1955).

According to our  $H_0$ , we postulate that there exists a normal frequency distribution of  $x = \log z$ . By  $H_0$  let

$$N[u(x); 0, 1] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{u(x)} e^{-x^2/2} dx.$$
 (3)

In this case and only in this case the function u(x) is linear and is

$$u = \frac{x - a}{\sigma},\tag{4}$$

where a is the mathematical expectation of x and  $\sigma$  is the standard deviation; in equation (4) we estimate a by  $\bar{x}$  and  $\sigma$  by  $\hat{s}$ , where  $\hat{s}$  is the unbiased statistic

$$\hat{s} = \sqrt{\frac{1}{n-1} \left[ \sum x_i^2 - \left( \frac{\sum x_i}{n} \right)^2 \right]}$$
 (5)

and n is the number of observations. At our  $H_0$  for logarithms z we have a straight line for the point u(x) on the co-ordinate plane (x, u). Let us take our logarithms of concentrations of element in an ascending order and designate the points of the rise of the empirical function by  $\tilde{u}_i$ . It is obvious that the points  $(x, \tilde{u}_i)$  near the straight line u(x) are on the plane (x, u) for normal distribution.

The word "near" we understand thus: The standard deviation of values  $\tilde{u}_i$  of the observed distribution is normal and is given as

$$\sigma_{\tilde{u}_i} \approx \frac{1}{g(u_i)} \sqrt{\frac{0.25 - \left[\phi(u_i)\right]^2}{n}}, \quad (6)$$

where  $g(u_i)$  is the probability density and  $\phi(u_i)$  is the area of normal distribution. The tables of the normal probability distribution are to be found in every handbook of statistics. Using (6), we can find the function  $\mathcal{E}(\sigma_{\tilde{u}_i})$ , and by means of  $\mathcal{E}(\sigma_{\tilde{u}_i})$  we shall establish the bounds of the confidence intervals for  $\tilde{u}_i$  with the confidence coefficient, p. Taking the confidence coefficient b, we state that all the points of observation which lie between the straight line u(x) and the curves  $\mathcal{E}(\sigma_{\tilde{u}_i})$  are those "near the straight line u(x)"; all points which lie outside  $\mathcal{E}(\sigma_{\tilde{u}_i})$  are "far from the straight line u(x)." If the points  $\tilde{u}_i$  lie near u(x), we accept  $H_0$  with the confidence coefficient p. If the points  $ilde{u}_i$  lie outside  $\mathcal{E}(\sigma_{ ilde{u}_i})$ , we reject  $H_0$  with the confidence coefficient 1 - p. If the points  $\tilde{u}_i$  lie near or on  $\mathcal{E}(\sigma_{\tilde{u}_i})$ , we can neither reject nor accept  $H_0$ .

In table 3 an example of calculations is given.

In table 4 the results of the calculations of  $\tilde{u}_i$ ,  $u_i$ , and  $\sigma \tilde{u}_i$  are given. In graphic form they are repeated in figure 1. From table 4 and figure 1 it was concluded that the lognormal  $H_0$  for Ga must be rejected because some points of observations lie outside  $\mathcal{E}(\sigma_{\tilde{u}_i})$ . The  $H_0$  can be accepted for Sc. Other distributions of the points (for Zr, Pb, La, V, Cr [?] in Canadian granites) are situated near 15 per cent confidence limits. For this reason we cannot reject the lognormal  $H_0$ ; but the 15 per cent confidence coefficient is very small for the empirical function—that is why one cannot seriously accept the lognormal  $H_0$ . The situation is such that our confidence limits can be reached by many skew distributions and, in particular, by lognormals (Rasumovsky, 1940; Ahrens, 1954), cubic roots (Chayes, 1954), and so on.

The fourth and last group of published data is K/Rb in igneous rocks and chondrites; Rb<sub>2</sub>O/Tl<sub>2</sub>O in potassium minerals;

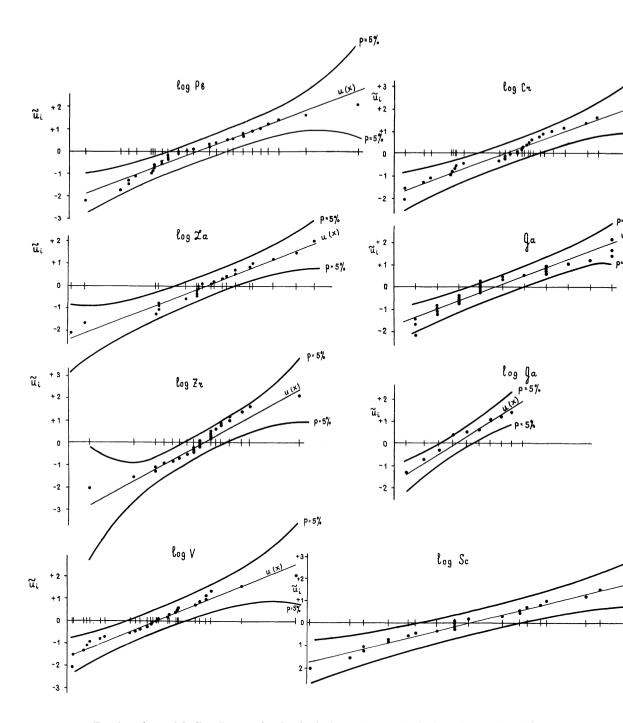


Fig. 1.—The straight-line diagrams for Ga, Cr, Sc, La, V, Pb, and Zn in Canadian granites. The dots are  $\tilde{u}_i$ ; curve  $\mathcal{E}(\sigma_{\tilde{u}_i})$ , are given for 5 per cent confidence limits  $(\rho)$  of  $\tilde{u}_i$ ; they are plotted from the line u(x) for the demonstration

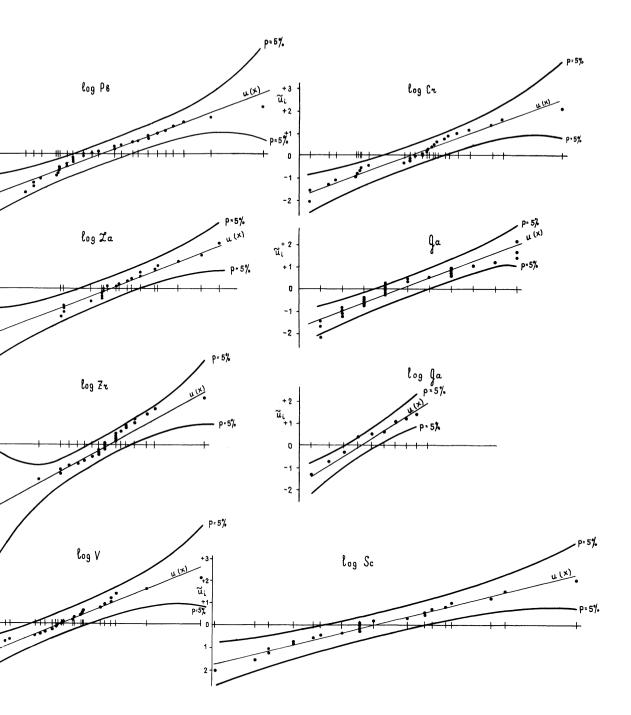


Fig. 1.—The straight-line diagrams for Ga, Cr, Sc, La, V, Pb, and Zn in Canadian granites. The dots are  $\tilde{u}_i$ ; curves,  $v_{\tilde{u}_i}$ ), are given for 5 per cent confidence limits  $(\rho)$  of  $\tilde{u}_i$ ; they are plotted from the line u(x) for the demonstration.

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No.	Sc	ln 10Sc	(100/n)(i-0.5)	$\tilde{u}_i + 5$	$ ilde{u}_i$	$X_i - \overline{X}$	$u_i = \frac{(X_i - \overline{X})}{\hat{s}}$	$\phi(u_i)$	$g(u_i)$	$\sqrt{0.25-[\phi(u_i)]^2}$	$\sqrt{n} \cdot g(u_i)$	$ \tilde{u}_i - u_i $	$\sigma_{ ilde{u}_i}$	$ \tilde{u}_i - u_i $ $1/\sigma_{\tilde{u}_i}$	$1.96\sigma_{\widetilde{u}_i}$
1	1.3	2.5649	2.174	2.980	-2.020	-1.6576	-1.754	0.4603	0.0857	0.1952	0.4110	0.266	0.475	0.560	0.931
2	2.0	2.9957	6.522	3.488	-1.512	-1.2268		0.4029	0.1718	0.2961	0.8239	0.214	0.359	0.596	0.704
3	2.3	3.1355	10.870	3.767	-1.233	-1.0870	-1.150	0.3749	0.2059	0.3308	0.9875	0.083	0.335	0.248	0.657
4	2.3	3.1355	15.218	3.973	-1.027	<b>-1.0870</b>		0.3749	0.2059	0.3308	0.9875	0.123	0.335	0.367	0.657
5	3.0	3.4012	19.566	4.143	-0.857	-0.8213		0.3076	0.2735	0.3942	1.3117	0.012	0.301	0.040	0.590
6	3.0	3.4012	23.913	4.290	-0.710	-0.8213		0.3076	0.2735	0.3942	1.3117	0.159	0.301	0.528	5.900
7	3.7	3.6109	28.261	4.425	-0.575	-0.6116		0.2412	0.3236	0.4379	1.5519	0.072	0.282	0.255	0.553
8	4.0	3.6889	32.609	4.549	-0.451	-0.5336		0.2140	0.3401	0.4519	1.6311	0.114	0.277	0.411	0.543
9	5.0	3.9120	36.957	4.667	-0.333	-0.3105		0.1289	0.3779	0.4831	1.8123	0.004	0.267	0.014	0.523
10	6.0	4.0943	41.305	4.780	-0.220	-0.1282		0.0541	0.3953	0.4971	1.8958	0.084	0.262	0.321	0.514
11	6.0	4.0943	45.653	4.890	-0.110	-0.1282		0.0541	0.3953	0.4971	1.8958	0.026	0.262	0.099	0.514
12	6.0	4.0943	50.000	5.000	0.000	-0.1282		0.0541	0.3953	0.4971	1.8958	0.136	0.262	0.519	0.514
13	6.0	4.0943	54.349	5.110	+0.110	-0.1282		0.0541	0.3953	0.4971	1.8958	0.246	0.262	0.939	0.514
14	7.0	4.2485	58.697	5.220	+0.220	+0.0260		0.0112	0.3988	0.4999	1.9126	0.192	0.261	0.736	0.512
15	10.0	4.6052	63.045	5.333	+0.333	+0.3827		0.1573	0.3675	0.4747	1.7625	0.072	0.269	0.268	0.527
16	12.0	4.7875	67.392	5.451	+0.451	+0.5650		0.2251	0.3336	0.4464	1.5999	0.147	0.279	0.527	0.547
17	12.0	4.7875	71.740	5.575	+0.575	+0.5650		0.2251	0.3336	0.4464	1.5999	$0.023 \\ 0.027$	0.279	0.082	0.547
18	13.0	4.8675	76.088	5.710	+0.710	+0.6450		0.2527	0.3159	0.4314	1.5150 1.3515	0.027	0.283	0.093	0.582
19	15.0	5.0106	80.436	5.857	+0.857	+0.7881		0.2978	0.2818 0.2656	0.4016 0.3870	1.3313	0.023	0.297	0.411	0.596
20	16.0	5.0752	84.784	6.027	+1.027	+0.8527		0.3165 0.4084	0.2050	0.3870	0.7889	0.123	0.366	0.411	0.390
21	24.0 28.0	5.4806	89.132	6.233	$\begin{array}{r} +1.233 \\ +1.512 \end{array}$	+1.2581		0.4084	0.1045	0.2508	0.7889	0.098	0.300	0.208	0.786
22 23	60.0	5.6348 6.3969	93 .480 97 .828	6.512 7.020		$\begin{array}{c c} +1.4123 \\ +2.1744 \end{array}$		0.4323	0.1303	0.1030	0.0239	0.017	0.759	0.370	1.684
20	00.0	0.3909	91.020	7.020	+2.020	T4.1744	<b>∓</b> 2.301	0.4093	0.0200	0.1000	0.1337	0.201	0.739	0.570	1.007

 $\bar{x} \approx 4.2225$ ;  $\hat{s} \approx 0.9449$ ;  $P(|x - \bar{x}| > 0.939 \ \sigma_{\tilde{u}}) \approx 0.35$ 

 $\begin{tabular}{ll} TABLE~4\\ Check~of~Lognormal~and~Normal~(for~Gallium)~H_0\\ \end{tabular}$ 

#### LOGNORMAL Ho

	La				Ga				Sc				Pb		
log La	$\tilde{u}_i$	$u_i$	$\sigma_{ ilde{u}_i}$	ln Ga	$ ilde{u}_i$	$u_i$	$\sigma_{\widetilde{u}_i}$	ln Sc	$\tilde{u}_i$	ui	$\sigma_{\widetilde{u}_i}$	ln Pb	$\tilde{u}_i$	$u_i$	$\sigma_{\widetilde{u}_i}$
. 1139	-2.020 -1.512 -1.233 -1.027 -0.857 -0.710 -0.575 -0.451 -0.333 -0.220 -0.110 0.000 +0.110 +0.220 +0.333 +0.451 +0.575 +0.710 +0.857 +1.027 +1.233 +1.512 +2.020	-2.307 -2.073 -0.833 -0.783 -0.783 -0.312 -0.136 -0.103 -0.039 +0.112 +0.168 +0.301 +0.538 +0.538 +0.538 +0.784 +1.187 +1.582 +1.899	0.613 0.297 0.293 0.293 0.293 0.272 0.264 0.264 0.264 0.263 0.264 0.267 0.277 0.277 0.277 0.277 0.303 0.345 0.425	2.5649 2.5649 2.5649 2.5649 2.6391 2.6391 2.7081 2.7081 2.7081 2.7081 2.7081 2.7726	-2.154 -1.675 -1:418 -1.230 -1.078 -0.946 -0.831 -0.724 -0.626 -0.533 -0.445 -0.360 -0.278 -0.118 -0.039 +0.118 +0.197 +0.278 +0.360 +0.445 +0.533 +0.626 +0.724 +0.831 +0.946 +1.078 +1.230 +1.418 +1.675 +2.154	-1.497 -1.497 -1.025 -1.025 -1.025 -0.587 -0.587 -0.587 -0.587 -0.178 -0	0.340 0.270 0.270 0.270 0.270 0.236 0.236 0.233 0.223 0.223 0.223 0.223 0.223 0.223 0.223 0.223 0.223 0.223 0.225 0.259 0.259 0.259	3.9120 4.0943 4.0943 4.0943 4.0943 4.2485 4.6052 4.7875 4.7875 5.0106 5.0752 5.4886 5.6348 6.3969	-2.020 -1.512 -1.233 -1.027 -0.859 -0.710 -0.575 -0.451 -0.333 -0.220 -0.110 0.000 +0.110 +0.220 +0.333 +0.451 +0.575 +0.710 +0.857 +1.027 +1.233 +1.512 +2.020	-1.754 -1.298 -1.150 -0.869 -0.869 -0.647 -0.565 -0.329 -0.136 -0.136 -0.136 +0.028 +0.405 +0.598 +0.683 +0.834 +0.902 +1.331 +1.495 +2.301	0.359 0.335 0.301 0.301 0.282 0.277 0.266 0.262 0.262 0.262 0.262 0.279 0.279 0.279 0.279 0.304 0.306 0.401	4 .0073 4 .0943 4 .0943 4 .1744 4 .3175 4 .3438 4 .3820 4 .4427 4 .4927 4 .4998 4 .4998 4 .4998 4 .6052 4 .6052 4 .7005 4 .7875 4 .9416 4 .9416	-2.14 -1.66 -1.40 -1.21 -1.06 -0.93 -0.81 -0.70 -0.60 -0.51 -0.42 -0.33 -0.25 -0.16 -0.08 +0.08 +0.16 +0.25 +0.33 +0.42 +0.51 +0.60 +0.70 +0.81 +0.93 +1.06 +1.21 +1.40 +1.66 +2.14	-1.837 -1.255 -1.118 -0.992 -0.766 -0.725 -0.665 -0.569 -0.479 -0.479 -0.313 -0.313 -0.163 -0.026 +0.216 +0.216 +0.216 +0.522 +0.613 +0.778 +0.778 +0.928 +1.061 +1.1364 +1.830 +2.706	0.303 0.285 0.275 0.270 0.251 0.244 0.239 0.235 0.235 0.225 0.225 0.227 0.227 0.227 0.252

9

				Lognorma	L Ho							Normal	<i>H</i> <sub>0</sub>	
Cr				Zr				v				Ga		
$\log$ Cr $ ilde{u}_i$	$u_i$	$\sigma_{\widetilde{u}_i}$	log Zr	ũi	$u_i$	$\sigma_{\widetilde{u}_i}$	log V	$ ilde{u}_i$	$u_i$	$\sigma_{\widetilde{u}_i}$	Ga	ũi	ui	$\sigma_{\widetilde{u}_i}$
0.3010.	3	0.276 0.274 0.265 0.242 0.241 0.241 0.243 0.243 0.244 0.245 0.255 0.259 0.262 0.287 0.348 0.378	1.9031 2.0000 2.0414 2.0792 2.1139 2.1461 2.1761 2.1761 2.2041 2.2041 2.2041 2.2041 2.2041 2.2041 2.2553	-2.086 -1.593 -1.325 -1.128 -0.967 -0.828 -0.704 -0.590 -0.483 -0.381 -0.282 -0.187 -0.093 +0.187 +0.2822 +0.381 +0.483 +0.590 +0.704 +0.828 +0.967 +1.128 +1.325 +1.593 +2.086	-1.673 -1.124 -0.890 -0.676 -0.479 -0.127 -0.127 +0.033 +0.033 +0.033 +0.325 +0.325	0.306 0.279 0.262 0.252 0.242 0.242 0.241 0.241 0.241 0.246 0.246 0.257 0.257 0.273 0.273 0.307 0.335	0.8451 0.8751 0.9031 1.0000 1.0414 1.2788 1.3222 1.3617 1.4314 1.4771 1.5185 1.5315 1.6232 1.6335 1.6990 1.7076 1.7160 1.7160 1.7160 1.7160 1.7160 1.7160 1.7160 1.7160 1.7151 1.9138 1.9731 1.9731 2.1584 2.3010	-2.086 -1.593 -1.325 -1.128 -0.967 -0.828 -0.704 -0.590 -0.483 -0.282 -0.187 -0.903 +0.187 +0.2822 +0.381 +0.483 +0.590 +0.704 +0.828 +0.828 +1.128 +1.325 +1.593 +2.086	-1.542 -1.527 -1.334 -1.218 -1.025 -0.943 -0.471 -0.384 -0.306 -0.167 -0.104 -0.076 +0.032 +0.235 +0.235 +0.388 +0.388 +0.388 +0.400 +0.716 +0.793 +0.911 +1.280 +1.564 +2.556	0.377 0.338 0.319 0.293 0.293 0.285 0.245 0.242 0.241 0.241 0.241 0.243 0.244 0.247 0.243 0.248 0.271 0.271 0.281 0.329 0.386	13	-2.154 -1.675 -1.418 -0.946 -0.831 -0.724 -0.626 -0.533 -0.445 -0.360 -0.287 -0.197 -0.118 -0.039 +0.118 +0.197 +0.287 +0.360 +0.445 +0.533 +0.626 +0.724 +0.831 +0.946 +1.078 +1.230 +1.418 +1.675 +2.154	-0.985 -0.985 -0.614 -0.614 -0.614 -0.243 -0.243 -0.243 -0.243 -0.243 +0.128 +0.128 +0.498 +0.869 +0.869 +1.240 +1.611 +1.981	0.314 0.314 0.266 0.266 0.266 0.266 0.237 0.237 0.237 0.224 0.224 0.224 0.224 0.224 0.225 0.255 0.255 0.250 0.480 0.480 0.480

Sr/Ca in limestones (Ahrens, 1954b); Ni/Fe, Zn/Fe, Cu/Fe, and Mn/Fe in pyrrhotites (Durovič, 1957). These frequency distributions are in contradiction to the lognormal  $H_0$  for K, Rb, Sr, etc., because we know no reason why their ratio with other elements should be lognormal. It must be admitted that the problem of the probability distribution of products of random variables and their ratios is very complicated, especially for correlated random variables. It can be illustrated by the simplest products of two variables, each of them having a normal distribution and being correlated with one another (Aroian, 1947). In our case we have, according to  $H_0$ , lognormally distributed random variables (Ahrens, 1954b, 1957) with a strong correlation (K-Tl, for example) between them. In this case we do not know the distribution function of the ratio. In the case of Sr/Ca and Durovič's example (1957) we have the minor elements Sr, Ni, Zn, Cu, Mn, and the major ones, such as Ca in limestones and Fe in pyrrhotites. For the minor elements we can propose the lognormal  $H_0$ , but for the major elements this cannot be done from any general considerations, as was pointed out by Aubrey (1956). It is obvious that if there are great quantities of the element in any minerals or rocks—like Fe in pyrrhotites the expectation of a great positive asymmetry is hardly real. New data on this question have not been published; therefore, we do not know what probability distributions will have ratios of minor to major elements of rocks or minerals.

Neither is it clear why F(z) must be lognormal if F(x) and F(y) are lognormal and

$$F(z) = F(x) *F(y)$$
.

where \*\* is the designation of the composition (Cramer, 1946), as was proposed by Rasumovsky for the distributions of Zn + Pb, every one of which has a lognormal distribution according to his hypothesis.

Our analysis of the published data is completed.

The lognormal distribution with a high confidence coefficient is a very rare case. The lognormal distribution can be rejected in very rare cases. In a great number of cases we can neither reject nor accept the lognormal  $H_0$ . The problem must be treated according to more sufficient data. This treatment will be given in the second part of this paper.

#### ANALYSIS OF SKEWED DISTRIBUTIONS

From the materials examined in the previous publications it follows that in many cases the frequency distributions of the concentrations of elements have positive skewness. These distributions can sometimes be sufficiently approximated by a lognormal function. Some of these distributions have a large skewness but are not lognormal. The fact is that the great positive skewness is widely observed and must be interpreted from the geochemical point of view. Two schemes can be given for the explanation of the skewness of distributions.

The first scheme has been cited by Miller and Goldberg (1955) taken from Cramer's handbook (1946). This scheme was proposed by Kaptein and Van Uven in the years 1903 and 1916. Its sense is the following: Let us have a concentration, X, which may be regarded as the sum of a large number of impulses  $z_1, \ldots, z_n$ . These impulses are mutually independent and act in the order of their subscripts. Let the increase produced by the impulse  $z_r$  be assumed to be directly proportional to  $z_r$  and to the concentration X. The concentration  $X_r$  results from the action of r - 1 of the preceding impulses.

In this case

$$X_{\nu+1}-X_{\nu}=k\,z_{\nu}X_{\nu}\,\,,$$

and hence

$$z_1 + \ldots + z_n = \frac{1}{k} \sum_{1}^{n} \frac{X_{\nu+1} - X_{\nu}}{X_{\nu}}.$$
 (7)

If this process is discontinuous,

$$w = \frac{1}{k} \int_{X_1}^{X_2} \frac{dt}{g(t)},$$
 (8)

and w will be normally distributed, so that we shall have a lognormal distribution. It is obvious from equation (8) that we shall get a lognormal distribution only when g(t) is a simple linear function, i.e., g(t) = at. In other cases we shall have various functions which would be very troublesome to use in normal distributions.

From the foregoing, we have only one population. This population is a lognormal population, and it occurs independently whether we have a whole batholith or a little bit of it.

The analyzed scheme assumes one frequency distribution function independently, whether we have sampled a piece of the rock or a geological body formed by this rock.

The second scheme is the following: The geochemical process at every separate moment gives a concentration of the elements in the conditions of the central-limit theorem (Cramer, 1946). In the course of time the mean values of the deposited concentrations increase. At the same time the values of variances also increase. The variability of the mean values and variances have a linear positive correlation as an approximate estimate of their relations. In other words, the probability density, f(x), is

$$f(x) = \sum_{i} p_{i} f_{i}(x), \qquad (9)$$

where  $f_i(x)$  is normal density,  $p_i$  is the function of weight of summarized densities, and i is the stage of the geochemical process put in the order of time.<sup>4</sup>

In this case we have a skewed function f(x), but this function, whether lognormal or not has only one certain feature—the positive skewness of f(x) when  $p_i$  is given the order of non-increasing values.

The principal difference of the last scheme from lognormal is the following: the frequency distribution functions for one separate stage of the geochemical process<sup>5</sup>

are normal. The frequency distribution functions for all joint products of the geochemical process have a positive skewness (we called these distributions "joint distributions").

The given schemes can be checked in an experimental way. For this purpose we must investigate the following questions: (a) whether the skewed distributions of the concentration of elements for the joint products of the geochemical process are lognormal and, if these distributions are not lognormal, whether we have a positive linear correlation between  $\bar{x}$  and s for these distributions, and (b) whether the positively skewed and not lognormal distributions of joint products of the geochemical process have normal distributions for fixed stages of this geochemical process.

## THE JOINT DISTRIBUTIONS, THEIR AVERAGES AND STANDARD DEVIATIONS

As an example of the joint distribution, we have investigated the distribution of phosphorus in granitic rocks. For this reason we compiled analytical data of P<sub>2</sub>O<sub>5</sub> in the granites of the world which contain more than 60 per cent SiO<sub>2</sub>. The frequency distribution for P2O5 in granitic rocks of Switzerland is given in figure 2. This distribution is valuable because the sampling was produced on a small territory and gives many analytical data. As we can see from figure 2, the analyzed distribution of P<sub>2</sub>O<sub>5</sub> has a great positive skewness but is not a lognormal one, as is indicated by the estimations of its statistics (shown in fig. 2). So the first given distribution of the element cited by Ahrens (1957, p. 206) has a positive skewness but is not lognormal.

Table 5 gives the values of the averages of  $P_2O_5$  and its standard deviations in some regions of the world. These figures are calculated from a number of analyses collected from many papers. We shall give the list of the papers used in our next article.

Proceeding as before (Vistelius, 1948, 1958), we got a coefficient of correlation be-

<sup>&</sup>lt;sup>4</sup> The process of growing of the mean values and variances can be continuous.

<sup>&</sup>lt;sup>5</sup> Let us name them "local distributions."

tween the average contents of  $P_2O_5$  in samples and their standard deviation as 0.56. The equation of regression of the standard deviation on averages is as follows:

$$s = 0.601 \overline{P_2 O_5} + 0.0538$$
.

From our analysis it is obvious that  $P_2O_5$  has a positive skew distribution and a positive correlation between the average values and standard deviations of its distributions.

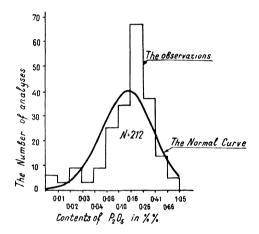


Fig. 2.—Frequency distribution of  $P_2O_5$  in the granitic rocks of Switzerland (with  $SiO_2 > 60$  per cent).

### JOINT DISTRIBUTIONS VERSUS LOCAL DISTRIBUTIONS

In figure 3 there is given a histogram of the frequency distributions of Na<sub>2</sub>O obtained by 4,788 analyses of igneous rocks (Washington, 1917). As shown in figure 3, the distribution of Na<sub>2</sub>O in igneous rocks has great positive skewness and is not a lognormal one. So the joint distribution of Na<sub>2</sub>O is a skew and not a lognormal one. In figure 4 is given the distribution of Na<sub>2</sub>O in the sample of 200 analyses of the basalts of the world. This distribution is also skewed, but the skewness is smaller than in the previous case.

From these examples the following conclusions can be drawn. All the rocks of the world are a collection of results of a very wide range of stages of the geochemical process. Their joint distribution is very skew. All the basalts of the world are also a collection of the results of the range of stages of the geochemical process, but their range of the geochemical process is narrower than the range of those of all rocks. If we turned our attention to the skewness of the analyzed distributions, we could obtain, without any calculations, the notion that the skewness of the joint distribution of all the rocks is greater than that of the joint distribution of basalts.

We shall formulate the next hypothesis from observations in the following way: local distributions are normal, whereas the joint distributions are skewed. We checked our hypothesis by the following procedure.

In table 6 are given the figures for Na<sub>2</sub>O concentrations in the basalts of the world. We placed in one row of table 6 the analyses with very similar contents of SiO<sub>2</sub>. They are given for basalts taken from one definite place. In the columns of table 6 are given the analyses of basalts from the different geographical locations and with different compositions. We have rocks varying from leucite basalts to nepheline-melilite rocks and from taitites to andesite basalts.

For the understanding of the local distributions we proceed as follows (Dunin-Barkovsky and Smirnov, 1955).

Let us designate the percentage of  $Na_2O$  in basalts (table 6) by x and let

$$\tau_i = \frac{x_i - \bar{x}}{\hat{s}},\tag{10}$$

where  $\bar{x}$  and  $\hat{s}$  are designated as previously. In the values of  $\tau$  we have a statistic which is free of the mean value of the distribution and its standard deviation but contains skewness and excess of distribution. We obtain values of  $\tau$  by sliding from the upper left corner of table 6 to the lower right corner by four rows and then again to the left side of the table.

According to Cramer (1946), we know that

$$\phi(\tau) = \frac{\Gamma[(n-1)/2]}{\sqrt{\pi(n-1)}\Gamma[(n-2)/2]} \cdot \left(1 - \frac{\tau^2}{n-1}\right)^{(n-4)/2} \quad (11)$$
for  $|\tau| < \sqrt{n-1}$ .

But we have taken n=4, so on the right side of equation (11) we have constants; consequently, the distribution of  $\tau$  in our case is uniform. Integrating equation (11) by  $\tau$ , we obtain theoretical values  $F_i$  which

are cumulative probability distribution functions of the normal  $H_0$  for values  $\tau_i$ ,

$$F_i = [\tau_i - (-\sqrt{3})] \cdot 0.289$$
.

We calculate the cumulative frequency function,  $\tilde{F}_{i}$ , by means of

$$\tilde{F}_i = \frac{i - 0.5}{n} \,. \tag{12}$$

The values of  $F_i$  and  $\tilde{F}_i$  and  $|F_i - \tilde{F}_i|$  are given in table 7. The probability  $P_n(\lambda)$ 

 $TABLE \ 5 \\$  Average Values and Standard Deviations  $P_2O_5$  in Samples from the Granitic Rocks with SiO2 More than 60 Per Cent\*

No. Region	$\overline{P_2O_5}$	$S_{P_2O_5}$	n
1 North Finland and Kola peninsul		0.188	50
2 South and central Finland (on the from Rovaniemi)	e south	0.178	141
3 Ukranian Shield	0.17	0.178	119
4 Middle Ural	0.14	0.101	48
5 Northwest Caucasus	0.14	0.099	93
6 Taimyr peninsula and North Land		0.123	43
7 Central Kazahstan and western A		0.127	173
8 Southwestern Tienshan	0.11	0.086	106
9 Anabara Shield	0.12	0.134	13
10 Watershed region Baikal Lake-Vi	tim		
River	0.15	0.142	50
11 Transbaikal	0.14	0.158	107
12 Northeast U.S.S.R. (from the lin	e Vitim-		
Lena Rivers)	0.08	0.070	51
13 On the Southeast of Amur-Ussuri		0.211	33
14 Japan	0.20	0.214	134
15 Eastern China	0.24	0.120	46
16 Indochina	0.12	0.087	42
17 East Indies Islands and Malacca	0.11	0.286	70
18 Australia	0.14	0.142	120
19 British Columbia, Washington, Or		0.068	51
20 Montana, Wyoming, Colorado 21 California	0.13 0.12	0.123	52
22 Cantornia 22 Canadian Shield	0.12	0.125 0.151	76 56
23 New England and Adirondacks	0.19	0.131	72
24 South America	0.08	0.168	42
25 Greenland	0.10	0.108	25
26 Scotland, North Ireland, and Rock		0.197	74
27 England and Wales	0.22	0.126	28
28 Portugal	0.40	0.227	26
29 Atlas System	0.25	0.164	49
30 Sahara	0.09	0.081	41
31 Southwestern Africa	0.14	0.079	54
32 Belgian Congo, Angola, Northern I	Rhodesia 0.15	0.174	52
33 South Africa	0.14	0.132	121
34 Sweden and Denmark	0.18	0.164	83
35 Western Germany	0.28	0.289	64
36 Switzerland	0.20	0.164	222
37 Eastern Germany, Czechoslovakia		0.152	84
38 Madagascar	0.17	0.211	69

<sup>\*</sup> Contents of P2O5 are taken from the papers, a list of which will be published later.

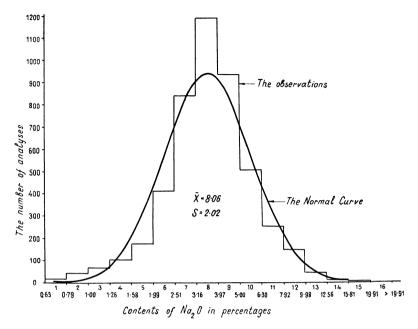


Fig. 3.—Frequency distribution of Na<sub>2</sub>O in igneous rocks of the world (logarithmic scale)

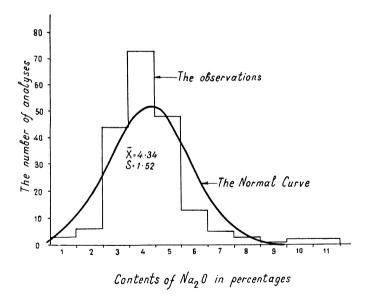


Fig. 4.—Frequency distribution of Na<sub>2</sub>O in the basalts of the world (arithmetic scale)

 $\label{eq:table 6} TABLE~6$  Contents of Na2O in Samples of Basalts

2 3	AUTHORS CITED  2  Melilite basalts, S. Africa (Hall,	X1	X2			X	ŝ	$X_i - \overline{X}$	$\tau_i = \frac{(X_i - X)}{\hat{S}}$
1 2 3 4				<i>X</i> 3	<i>X</i> <sub>4</sub>				$(A_i - X)/S$
2 3	Melilite basalts S Africa (Hall	3	4	5	6	7	8	9	10
4	1938)	2.42	2.03	3.48	3.95	2.970	0.8955	-0.5500	-0.6141
1	Basalts, S. Rhodesia (Hall, 1938) Alkali basalts, Kenya (Hall, 1938)	2.38	1.66 2.74		1.94 3.05		0.3352 0.3078	-0.2550	-0.7607
5	Basalts, Madagascar (Hall,	1.88					i	-0.3880	-1.2605
	1938) Basalts, Indochina (Lacroix, 1933)	3.15	2.01 3.34				j	+0.3400 $+0.0100$	+1.3077 $+0.0343$
6	Andesite basalts, North China (Lacroix, 1933)	3.31	3.28				0.0173	-0.0050	-0.2890
1	Leucitic basalts, Central Java (Van Bemmelen, 1937)	0.73	2.17				1.6773	-1.6320	-0.2890 $-0.9730$
8	Basalts, Aleutian Islands (Coats, 1952)	3.68	3.57	3.65	3.35	3.562	0.1641	+0.0080	+0.0487
	Plagioclase basalts, Sumatra (Westerveld, 1952)	3.43	3.29	3.19		3.345	0.1291	-0.1550	-1.2006
	Basalts, Krakatau group (Westerveld, 1952) Basalts, Armenia (Struve, 1940)	2.94 1.80	2.91 3.15				0.2754 0.7514	-0.2320	-0.8424
	Basalts, Chatanga, Jakutia (Struve, 1940)	2.24	1.91				0.1740	+0.6850	+0.9250
13	Basalts, Kamchatka (Struve, 1940)	2.82	2.15				}	-0.1150 +0.5550	-0.6609 $+1.2619$
14	Basalts, Hooker Island, Land of Frantz Josef (Struve, 1940)	2.31	1.80					+0.1300	+0.2058
15	Alkali basalts, Ussury River, far east of U.S.S.R. (Tatarinov,						ļ	·	,
16	1954) Basaltic limburgites, Hilok	4.03	2.68				0.5969	-0.3150	-0.5277
17	River, U.S.S.R. (Below, 1956) Basalts, Lower Silesia (Jerzman-	3.86						+0.2320	+0.3833
18	sky, 1956) Basalts, Australia (Yates, 1954)	3.27 2.72	2.70 2.98	4.04	3.68	3.355	0.6107	-0.5800 $-0.3750$	-1.0605 $-0.6142$
19 20	Basalts, Slovakyan (Salat, 1955) Basalts, Paricutin, Mexico	2.28			2.76			-0.6650	-1.2302
21	(Wilcox, 1954) Basalts, Armenia, U.S.S.R.	3.88	3.87					+0.0380	+0.3958
22	(Zavaritzky, 1953) Basalts, Avacha, Kamchatka	2.08	2.16		2.53	2.578	0.6683	+0.9620	+1.4394
23	(Gonshakova, 1953) Basalts, Fuji-San, Japan	3.18	2.83	2.45				+0.1200	+1.3886
24	(Tsuya, 1955) Basalts, Hungary (Mauritz and	l		2.78			0.1331	+0.1400	
25	Harwood, 1936) Basalts, Hungary (Jugovics,	1						·	+1.4088
26	1937) Basalts Réunion Island	1				i	l		-0.5623
27	(Lacroix, 1939) Ophitic basalts, California	1	1.29						+0.6498
28	(Powers, 1932) Olivine basalts, Juan Fernández	į					1		-0.6764
29	Island (Lacroix, 1928) Basalts, Mangareva Island (Lacroix, 1928)	1	2.94				0.5526 0.1222		+1.3427 $+0.4092$

TABLE 6-Continued

No.	Rocks, Localities, and Authors Cited	Con	TENTS O		) in	X	ŝ	$X_i - \overline{X}$	$ au_i =$
	AUTHORS CITED	<i>X</i> <sub>1</sub>	X 2	Хз	X4				$(X_i - \overline{X})/\hat{S}$
1	2	3	4	5	6	7	8	9	10
30	Taityte, Taity (Lacroix, 1928) Basalts, Mauna Loa, historic age	7.83			į		0.9145		
32	(Macdonald, 1948) Basalts, Kilauea, prehsitoric	2.43	2.30	2.07		2.272		+0.1580	
33	age (Macdonald, 1948) Basalts, Kilauea, historic age	2.01	2.42	2.20		2.198		+0.2220	
34	(Macdonald, 1948) Olivine basalts, Hualalai	2.64	2.20	2.35		2.410		-0.0600	
35	(Macdonald, 1948) Basalts, Mauna Kea	2.30			2.89		0.5401	-0.0950	
36	(Macdonald, 1948) Basalts, Kohala (Macdonald,	2.55		2.80		2.748	0.2425	+0.0520	
37	1948) Basalts, Faeroe Islands (Walker		3.56			3.075		+0.4850	,
38	and Davidson, 1936) Plagioclase basalts, the Azores	2.22	2.14			2.688		-0.4680	
39	(Essenwein, 1929/30) Olivine basalts, the Azores	3.54					0.4396	-0.2700	·
40	(Essenwein, 1929/30) Basalts, Spitsbergen (Tyrrelland	2.35		2.20			0.9701	-0.5220	•
41	Sanford, 1932/33) Nepheline-melilite basalts,	2.49	2.62	2.31			0.3756		,
42	Oahu (Winchell, 1947) Trachybasalts, Jan Mayen	3.93				4.585		-0.0450	
43	(Tyrrell, 1925/26) Basalts, Drakenberg, S. Africa	3.18	2.37	3.18	3.99	3.180	0.6614	+0.8100	-1.2247
44	(Walker and Poldervaart, 1949)	1.89	1.74	1.63	1.59	1.712	0.1424	+0.1780	+1.2500
44	Doleritic basalts, S. Africa, (Walker and Poldervaart, 1949)	1.92	1 8/	1 32	1 80	1 742	0.2878	+0.0980	+0.3405
45	Basalts, Mascarene Island (Walker and Nicolaisen,	1.92	1.01	1.02	1.07	1.772	0.2010	10.0000	10.0400
46	1954) Olivine basalts (Walker and	2.93	2.48	2.74	3.09	2.810	0.2626	-0.0700	-0.2665
47	Nicolaisen, 1954) Doleritic and olivine basalts	3.20	2.25	2.46	2.99	2.725	0.4441	+0.2650	+0.5967
27	(Walker and Nicolaisen, 1954)	2.90	3.08	2.77	2 97	2.930	0.1301	-0.1600	-1.2298
48	Trachybasalts, Tristan da Cunha and Gough Island	2.50	3.00	2,	2.71	2.550	3.1031	0.1300	1.22/0
49	(Smith, 1930)	5.84 1.84		7.28 2.47	5.25 2.87	5.845 2.300		-0.8350 $-0.4600$	
50	Basalts, West Greenland Noe-Nygaard, 1942)	0.39		1.01					

		~	,-		.~
i	Ti	$\widetilde{F}_i$	$\tau_i + \sqrt{3}$	$F_i$	$ \widetilde{F}_i - F_i $
1	2	3	4	5	6
1	-1.3253	0.010	+0.4068	+0.117	0.107
2	-1.2605	0.030	+0.4716	+0.136	0.106
3	-1.2301	0.050	+0.5020	+0.145	0.095
4	-1.2298	0.070	+0.5023	+0.145	0.075
5	-1.2247	0.090	+0.5074	+0.146	0.056
6	-1.2006	0.110	+0.5315	+0.153	0.043
7	-1.0605	0.130	+0.6716	+0.194	0.064
8	-0.9929	0.150	+0.7392	+0.213	0.063
9	-0.9730	0.170	+0.7591	+0.219	0.049
10	-0.8424	0.190	+0.8897	+0.257	0.067
11	-0.8201	0.210	+0.9120	+0.263	0.053
12	-0.7607	0.230	+0.9714	+0.280	0.050
13	-0.6764	0.250	+1.0557	+0.305	0.055
14	-0.0669 $-0.6142$	0.270 0.290	+1.0712	+0.309	0.039
15	-0.6142 $-0.6142$	0.290	$+1.1179 \\ +1.1179$	+0.323	0.033
16 17	-0.6142 $-0.6141$	0.310	+1.1179 +1.1180	$+0.323 \\ +0.323$	0.013 0.007
18	-0.5748	0.350	+1.1160 +1.1573	+0.323 +0.334	0.007
19	-0.5623	0.370	+1.1698	+0.334 +0.338	0.032
20	-0.5025	0.390	+1.2044	+0.348	0.032
21	-0.3247	0.410	+1.4074	+0.346	0.004
22	-0.2890	0.430	+1.4431	+0.417	0.013
23	-0.2665	0.450	-1.4656	+0.423	0.027
24	-0.1759	0.470	+1.5562	+0.449	0.021
25	-0.0903	0.490	+1.6418	+0.474	0.016
26	+0.0343	0.510	+1.7664	+0.510	0.000
27	+0.0487	0.530	+1.7808	+0.514	0.016
28	+0.2058	0.550	+1.9379	+0.559	0.009
29	+0.2144	0.570	+1.9465	+0.562	0.008
30	+0.3405	0.590	+2.0726	+0.598	0.008
31	+0.3833	0.610	+2.1154	+0.611	0.002
32	$+0.3886 \\ +0.3958$	0.630 0.650	$+2.1207 \\ +2.1279$	+0.612	0.018
33	+0.3938 +0.4092	0.670	+2.1279 +2.1413	$+0.614 \\ +0.618$	0.036 0.052
35	+0.4092 +0.4275	0.690	+2.1413 +2.1596	+0.623	0.032
36	+0.5381	0.710	+2.2702	+0.655	0.055
37	+0.5967	0.730	+2.3288	+0.672	0.058
38	+0.6311	0.750	+2.3632	+0.682	0.068
39	+0.6497	0.770	+2.3818	+0.688	0.082
40	+0.9250	0.790	+2.6571	+0.767	0.023
41	+0.9931	0.810	+2.7252	+0.787	0.023
42	+1.0518	0.830	+2.7839	+0.804	0.026
43	+1.2500	0.850	+2.9821	+0.861	0.011
44	+1.2619 +1.3077	0.870 0.890	+2.9940	+0.864	0.006
45	+1.3077 +1.3427	0.890	$+3.0398 \\ +3.0748$	+0.878	0.012
46	+1.3427 +1.3840	0.910	+3.0748	$+0.888 \\ +0.900$	0.022 0.030
48	+1.4088	0.950	+3.1101	+0.900 +0.907	0.030
49	+1.4110	0.930	+3.1431	+0.907	0.043
50	+1.4394	0.990	+3.1715	+0.916	0.074
			,	, 5.5.25	

 $\lambda_{50\%} = 0.8280 \; ; \quad \frac{\lambda_{50\%}}{\sqrt{50}} \approx 0.117 \; ; \quad p(|\tilde{F}_j - F_j| > 0.117) > 50\%$ 

is approximately equal to  $K(\lambda)$  for the inequality

$$\max |\tilde{F}_n - F| < \frac{\lambda}{\sqrt{n}} \qquad (13)$$

and for any  $\lambda$ , n > 40,

$$K(\lambda) = 1 - 2\sum_{\nu=1}^{\infty} (-1)^{\nu-1} e^{2\nu^2/2}.$$
 (14)

The values of  $\lambda_p$  for some levels of confidence have been tabulated (Dunin-Barkovsky and Smirnov, 1955). We give them in table 8.

The values  $\lambda_p$  which are given in table 8 show that normal  $H_0$  is in agreement with our observation, having a level of confidence of more than 50 per cent, as can be checked by tables 6, 7, and 8 and figure 5.

In this way the question is settled. The

TABLE 8

### Short Table of Percentile Deviations for Function $K(\lambda)$

(By Dunin-Barkovsky and Smirnov)

	Per	Per	Per	Per	Per
	Cent	Cent	Cent	Cent	Cent
<i>p</i>	50	10	5	1	0.1
$\lambda_p \dots$	0.828	1.224	1.358	1.627	1.950

frequencies of sodium are distributed normally with a high confidence level for fixed stages of the differentiation of the basalts.

#### THE LOCAL DISTRIBUTION

The best material which can be given for the study of the local distribution is a set of the chemical analyses of a small piece of rock. For this reason we took a piece of biotite granodiorite weighing nearly 1 kg., crushed it into 29 small bits of 5 gm. each, and determined with the greatest care the contents of  $P_2O_5$  in them.

The analyzed biotite granodiorite was taken by the author from the outcrop on the right slope of the gorge of the Pjandge River (Pimir) between Pasthoof and Barjumch villages.

The method of determining the  $P_2O_5$  is as follows (Dittler, 1933):

A piece of rock weighing nearly 5 gm. is crushed in an agate mortar and screened through a 100-mesh sieve. The sieved rock is melted with soda in a platinum crucible. The alloy is leached with warm water, filtered off of the insoluble remains, and then these remains are washed by means of a 1 per cent solution of soda. The filtrate is neutralized with nitric acid up to a weak acid reaction and is evapo-

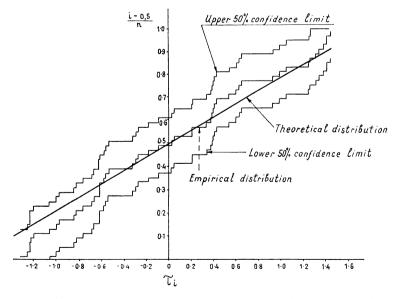


Fig. 5.—Cumulative frequency polygon, its probability distribution function and confidence limits

rated to 50–60 ml. To the solution are added 10 ml. of concentrated nitric acid and 20 ml. of ammonium nitrate (340 gm. NH<sub>4</sub>NO<sub>3</sub> in 11 H<sub>2</sub>O). The solution is heated up to 50° C., and 25 ml. of (NH<sub>4</sub>)<sub>2</sub>M<sub>0</sub>O<sub>4</sub>, heated up to 50° C., is added to this solution in the form of a thin squirt, while stirring it with a glass rod (the latter must not touch the walls of the glass). The yellow precipitate of ammonium phosphomolybdate soon separates out.

Identical conditions are strictly kept for precipitations of ammonium phosphomolybdate in all samples (volumes of solutions, their temperatures, the added quantity of reagents, etc.).

When the precipitation is completed, the glass is covered with a watch glass and is left overnight. The next morning the precipitate is filtered through a small dense filter, not shifting the precipitate from the glass to the filter. The precipitate is washed with the help of decantation and then with warm 5 per cent NH<sub>4</sub>NO<sub>3</sub> on the filter. The filtrate is mixed and left for the sake of checking the completeness of the precipitation.

Repeated analyses of one and the same piece of rock are given in table 9. From table 9 we see that the errors of determination of  $P_2O_5$  are not greater than 0.01 per cent.

The results of the analysis of the examined 29 pieces of granodiorite are given in table 10.

From figure 6 based on the data of table 10 it is obvious that the local distribution of  $P_2O_5$  in the analyzed biotite granodiorite is normal.

### THE FUNDAMENTAL LAW OF THE GEOCHEMICAL PROCESSES

We have stated that the probability distributions of the concentrations of chemical elements in the earth's crust are divided into two groups. The first group of distributions is that of the fixed stage of the geochemical process. These distributions are normal. The second group of distributions is that of mixed products of many stages of the geochemical process. The latter dis-

 $\label{thm:constraint} TABLE \ 9$  Determinations of  $P_2O_5$  Repeated in One and the Same Sample (Analyzer, Mrs. N. A. German)

	No	o. of the Pieces of	GRANODIORITE
	A	В	С
Determinations of P <sub>2</sub> O <sub>5</sub> repeated for one piece	0.300; 0.302	0.311; 0.317	0.316; 0.312; 0.314; 0.315

The washed precipitated ammonium phosphomolybdate is dissolved on the filter with warm ammoniac (1:2). The filter is washed several times in warm water with a small amount of ammoniac, and all the water used is collected, together with the solution. The solution and water are collected in an annealed and weighed-up porcelain cup (5.5 cm. in diameter), they are evaporated on a water bath, dried, and carefully annealed with a burner. In the course of annealing, the precipitate changes from white at the beginning to bright yellow and to dark blue at the end of the process (P<sub>2</sub>O<sub>5</sub>·24MoO<sub>4</sub>). The annealing must be very carefully made up to a temperature not exceeding 400° C. The annealing is ready when all the precipitate becomes dark blue. The cup containing the precipitate is cooled in an exsiccator and then weighed.

# TABLE 10 CONTENTS OF P<sub>2</sub>O<sub>5</sub> IN SMALL PIECES OF GRANODIORITE FROM PAMIR (Analyzer, Mrs. N. A. German)

	Content of		Content of
No. of	$P_2O_5$	No. of	$P_2O_5$
Pieces	(Per Cent)	Pieces	(Per Cent)
1	0.27	16	0.34
2	0.29	17	0.34
3	0.29	18	0.34
4	0.30	19	0.34
5	0.30	20	0.35
6	0.31	21	0.35
7	0.31	22	0.36
8	0.31	23	0.36
9	0.31	24	0.36
10	0.31	25	0.36
11	0.32	26	0.37
12	0.32	27	0.38
13	0.32	28	0.38
14	0.33	29	0.39
15	0.33		• • • •

tributions are skew, with a large positive skewness. Our analysis of geochemical peculiarities, by means of which we obtain the local and joint distributions, gives us the possibility of finding the most general features of the geochemical processes. We have named these features the "fundamental law of the geochemical processes," the fundamental law of the geochemical processes being as follows: The joint probability distribution function of the concentration of the minor chemical element deposited by natural chemical reactions has a large posi-

importance and has no relation to Gold-schmidt's point of view. The fact is that when Goldschmidt (1954) wrote: "Modern geochemistry studies the distribution and amount of the chemical elements in minerals, ores, rocks, soils, waters, and the atmosphere, and the circulation of the elements in nature, on the basis of the properties of their atoms and ions" he, as well as all other classicists of geochemistry such as V. I. Vernadsky, F. E. Clarke, A. E. Fersman, meant a deterministic approach to nature. In the cited papers published by

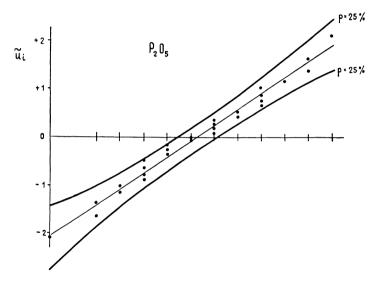


Fig. 6.—The straight-line diagram and confidence limits for the distribution of  $P_2O_\delta$  in the piece of granodiorite from Pamir.

tive skewness. This skewness indicates that the deposition of small concentrations of the minor element by geochemical processes is, as a rule, more stable than the deposition of large concentrations of this element by the same geochemical process—the quantity of the deposited element in the form of small concentrations being not less than that in the form of large concentrations. It seems to us that prospecting and mining confirm our statements.

We assume that the fundamental law of the geochemical processes has independent

<sup>6</sup>We did not include major elements or mechanical deposits.

Ahrens as well as the papers of the author, including the present one, we have the probabilistic approach to nature. Goldschmidt and Vernadsky spoke about concrete concentrations of the chemical elements, whereas we speak about probability distributions of the concentrations. These are quite different things. For instance, let us take two flakes of mica. In the first flake we have 0.1 per cent Rb<sub>2</sub>O, in the second one we have 0.2 per cent Rb<sub>2</sub>O, with normal probability distributions in both flakes. From Goldschmidt's point of view, the distributions of Rb<sub>2</sub>O<sub>0</sub> in the flakes are different because in one of them the concentration of Rb<sub>2</sub>O is

twice as great as that in the second. From our point of view, the analyzed flakes have similar normal probability distributions of the concentrations of Rb<sub>2</sub>O. Further, let us take two pieces of feldspar, one of which has a normal and the second a lognormal probability distribution of the concentration of Cs<sub>2</sub>O, with average concentrations of Cs<sub>2</sub>O in every piece equaling 0.01 per cent; From Goldschmidt's point of view the pieces have equal distributions of Cs<sub>2</sub>O, because in every piece it equals 0.01 per cent; from our point of view the probability distributions in the analyzed pieces are different because one is normal and the other is lognormal; by means of the analysis of the stochastic sense of the probability distributions of the concentrations of Cs2O in pieces, we can understand the difference between the processes of the concentration of Cs<sub>2</sub>O in the minerals of the pieces analyzed.

In modern science we deal with the increasing role of the probabilistic approach to nature. The formulation of the fundamental law of geochemical processes is the reflection of modern tendencies in science.

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